

Chapter 6A: Moments, Skewness, and Kurtosis

MATH/STAT 394: Probability I

Arman Jahangiri

University of Washington Department of Mathematics

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By the end of this lecture, students should be able to:

- ▶ define and interpret the mean, median, and mode;
- ▶ define moments, standardized moments, and standardized moments;
- ▶ define skewness and kurtosis;
- ▶ understand symmetry of distributions;
- ▶ define sample moments;
- ▶ compute the mean and variance of the sample mean.

Summaries of a Distribution

A probability distribution contains a large amount of information.

Often we want simpler numerical summaries:

- ▶ Where is the center?
- ▶ How spread out is the distribution?
- ▶ Is it symmetric?
- ▶ Are extreme values common?

Main idea

Moments provide numerical summaries of distributions.

Three common summaries of the center are:

- ▶ Mean (expected value)
- ▶ Median
- ▶ Mode

These summarize different aspects of a distribution.

Important

Different notions of “center” can give very different values.

Definition

A number m is a median of X if

$$\mathbb{P}(X \leq m) \geq \frac{1}{2}$$

and

$$\mathbb{P}(X \geq m) \geq \frac{1}{2}.$$

Intuitively:

- ▶ at least half the distribution lies to the left;
- ▶ at least half lies to the right.

Medians are less sensitive to extreme values than means.

Definition

A value c is called a *mode* of a random variable X if it maximizes:

$$\begin{cases} p(c), & \text{for a discrete random variable with PMF } p(x), \\ f(c), & \text{for a continuous random variable with PDF } f(x). \end{cases}$$

Equivalently,

$$\begin{cases} p(c) \geq p(x), & \forall x, \\ f(c) \geq f(x), & \forall x. \end{cases}$$

Interpretation

The mode is the most likely value (discrete case) or the point of highest density (continuous case).

Suppose salaries are:

40, 42, 45, 46, 48, 50, 1000.

Then:

median = 46,

while

mean $\approx$ 181.6.

Observation

The mean is heavily affected by extreme values (outliers), while the median is robust.

Best Prediction Depends on Loss Function

Suppose we predict an unknown random variable X by a constant c .

Two common error measures:

$$\mathbb{E}(X - c)^2$$

(mean squared error)

and

$$\mathbb{E}|X - c|$$

(mean absolute error).

Key fact

The mean minimizes mean squared error.

A median minimizes mean absolute error.

Mean and Variance Are Not Enough

Two distributions can have:

same mean, same variance,

yet look very different.

Possible differences:

- ▶ asymmetry;
- ▶ heavy tails;
- ▶ sharp peaks;
- ▶ multiple modes.

Conclusion

Higher moments help describe these additional features.

Moments and Symmetry

Definitions

The n th population moment:

$$\mu'_n = \mathbb{E}(X^n).$$

The n th central population moment:

$$\mu_n = E[(X - \mu)^n]$$

The n th standardized population moment:

$$\tilde{\mu}_n = \frac{\mu_n}{\sigma^n} = \mathbb{E}\left(\left(\frac{X - \mu}{\sigma}\right)^n\right),$$

So, we have the following relations

- ▶ **Mean** is the 1st population moment $\mu'_1 = \mu$, determines center
- ▶ **Variance** is the 2nd standardized population moment $\mu_2 = \sigma^2$, determines spread
- ▶ **Skewness** is the 3rd standardized population moment $\frac{\mu_3}{\sigma^3}$, determines asymmetry
- ▶ **Kurtosis** is the 4th standardized population moment $\frac{\mu_4}{\sigma^4}$, determines tail behavior

Definition

A random variable X is symmetric about μ if

$$X - \mu \stackrel{d}{=} \mu - X = -(X - \mu),$$

where $\stackrel{d}{=}$ denotes equality in distribution.

Let

$$Y = X - \mu.$$

Then the CDF of Y is

$$F_Y(y) = P(Y \leq y) = P(X - \mu \leq y) = P(X \leq y + \mu) = F_X(y + \mu).$$

If X has PDF f_X , then differentiating gives

$$f_Y(y) = f_X(y + \mu).$$

Similarly, for $-Y = \mu - X$,

$$f_{-Y}(y) = f_Y(-y) = f_X(\mu - y).$$

Symmetry About a Point (Continue)

Therefore, X is symmetric about μ when

$$f_X(\mu + y) = f_X(\mu - y) \quad \text{for all } y.$$

- ▶ Values equally far to the left and right of μ have the same density.
- ▶ The graph of f_X is a mirror image across the vertical line $x = \mu$.
- ▶ If the mean exists, then the center of symmetry is the mean:

$$E[X] = \mu.$$

- ▶ More generally, all odd central moments that exist are zero:

$$E[(X - \mu)^{2k+1}] = 0.$$

Symmetry and Odd Central Moments

Suppose X is symmetric about μ . Then

$$f_X(\mu + t) = f_X(\mu - t).$$

Consider an odd central moment, say the $(2k + 1)$ st:

$$\mu_{2k+1} = \mathbb{E}[(X - \mu)^{2k+1}] = \int_{-\infty}^{\infty} (x - \mu)^{2k+1} f_X(x) dx.$$

Make the change of variable

$$t = x - \mu.$$

Then $x = \mu + t$, so

$$\mu_{2k+1} = \int_{-\infty}^{\infty} t^{2k+1} f_X(\mu + t) dt.$$

Now define

$$g(t) = t^{2k+1} f_X(\mu + t).$$

Then

$$\begin{aligned} g(-t) &= \underbrace{(-t)^{2k+1}}_{=-t^{2k+1}} \underbrace{f_X(\mu - t)}_{=f_X(\mu+t)} \\ &= -g(t). \end{aligned}$$

Thus g is an **odd function**. Therefore,

$$\underbrace{\int_{-\infty}^{\infty} g(t) dt}_{\text{integral of an odd function over a symmetric interval}} = 0.$$

Hence

Proposition

Let X be a random variable symmetric about μ . Then, for every integer $k \geq 0$,

$$\mathbb{E}[(X - \mu)^{2k+1}] = 0,$$

provided the moment exists.

Proof.

Let $Y = X - \mu$. By symmetry, $Y \stackrel{d}{=} -Y$. Therefore, $\mathbb{E}[Y^{2k+1}] = \mathbb{E}[(-Y)^{2k+1}]$. But since $2k + 1$ is odd, $(-Y)^{2k+1} = -Y^{2k+1}$. Hence, $\mathbb{E}[Y^{2k+1}] = -\mathbb{E}[Y^{2k+1}]$. So, $2\mathbb{E}[Y^{2k+1}] = 0$, and therefore $\mathbb{E}[Y^{2k+1}] = 0$. Since $Y = X - \mu$, $\mathbb{E}[(X - \mu)^{2k+1}] = 0$. □

Examples of Symmetry

A distribution is symmetric about μ when points equally far from μ have the same probability density:

$$f_X(\mu + t) = f_X(\mu - t).$$

- **Normal distribution:** If $X \sim N(\mu, \sigma^2)$, then

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right).$$

For any t ,

$$f_X(\mu + t) = \frac{1}{\sigma\sqrt{2\pi}} e^{-t^2/(2\sigma^2)} = f_X(\mu - t).$$

Thus, the normal distribution is symmetric about μ .

- **Uniform distribution on a symmetric interval:** If $X \sim \text{Unif}(\mu - a, \mu + a)$, then

$$f_X(x) = \begin{cases} \frac{1}{2a}, & \mu - a \leq x \leq \mu + a, \\ 0, & \text{otherwise.} \end{cases}$$

Hence, for every t , $f_X(\mu + t) = f_X(\mu - t)$, since either both points lie inside the interval and have density $\frac{1}{2a}$, or both lie outside and have density 0.

- **Binomial distribution with $p = \frac{1}{2}$:** If $X \sim \text{Bin}(n, \frac{1}{2})$, then

$$P(X = k) = \binom{n}{k} \left(\frac{1}{2}\right)^n = \binom{n}{n-k} \left(\frac{1}{2}\right)^n = P(X = n - k).$$

Thus, the distribution is symmetric about $\frac{n}{2}$.

Definition

The skewness of X is

$$\text{Skew}(X) = \mathbb{E} \left(\left(\frac{X - \mu}{\sigma} \right)^3 \right).$$

Interpretation:

- ▶ positive skewness: long right tail;
- ▶ negative skewness: long left tail;
- ▶ zero skewness: often indicates symmetry.

Examples of Skewness

- ▶ Exponential distributions are right-skewed.
- ▶ Income distributions are often right-skewed.
- ▶ Normal distributions have skewness 0.

Important

Zero skewness does not necessarily imply symmetry.

Definition

The kurtosis of X is

$$\text{Kurt}(X) = \mathbb{E} \left(\left(\frac{X - \mu}{\sigma} \right)^4 \right).$$

- ▶ Kurtosis measures tail heaviness and extremeness.
- ▶ $\text{Kurt}(N(\mu, \sigma^2))=3$. If $\text{Kurt}(X) > < 3$ then f_X has heavier/lighter tails compared to normal.

Common Discrete Distributions

Distribution	Mean	Median	Mode	Variance	Skewness	Kurtosis
Bernoulli(p)	p	$\begin{cases} 0, & p < 1/2 \\ 1/2, & p = 1/2 \\ 1, & p > 1/2 \end{cases}$	$\begin{cases} 0, & p < 1/2 \\ 0, 1, & p = 1/2 \\ 1, & p > 1/2 \end{cases}$	$p(1-p)$	$\frac{1-2p}{\sqrt{p(1-p)}}$	$\frac{1-6p(1-p)}{p(1-p)} + 3$
Binomial(n, p)	np	$\approx np$	$\lfloor (n+1)p \rfloor$	$np(1-p)$	$\frac{1-2p}{\sqrt{np(1-p)}}$	$\frac{1-6p(1-p)}{np(1-p)} + 3$
Discrete Uniform($1, \dots, n$)	$\frac{n+1}{2}$	$\frac{n+1}{2}$	none	$\frac{n^2-1}{12}$	0	$\frac{9}{5} \cdot \frac{n^2+1}{n^2-1}$
Geometric(p)	$\frac{1}{p}$	$\left\lceil \frac{-1}{\log_2(1-p)} \right\rceil$	1	$\frac{1-p}{p^2}$	$\frac{2-p}{\sqrt{1-p}}$	$6 + \frac{p^2}{1-p} + 3$
Poisson(λ)	λ	$\approx \lfloor \lambda + \frac{1}{3} - \frac{0.02}{\lambda} \rfloor$	$\lfloor \lambda \rfloor - 1, \lfloor \lambda \rfloor$	λ	$\frac{1}{\sqrt{\lambda}}$	$\frac{1}{\lambda} + 3$

Click [here](#) to see a bigger table, which you can bring a copy on the exams too. It is on the instructor's course website.

Common Continuous Distributions

Distribution	Mean	Median	Mode	Variance	Skewness	Kurtosis
Uniform(a, b)	$\frac{a+b}{2}$	$\frac{a+b}{2}$	none	$\frac{(b-a)^2}{12}$	0	$\frac{9}{5}$
Normal(μ, σ^2)	μ	μ	μ	σ^2	0	3
Exponential(λ)	$\frac{1}{\lambda}$	$\frac{\log 2}{\lambda}$	0	$\frac{1}{\lambda^2}$	2	9
Gamma(α, λ)	$\frac{\alpha}{\lambda}$	(no closed form)	$\frac{\alpha-1}{\lambda} \quad (\alpha > 1)$	$\frac{\alpha}{\lambda^2}$	$\frac{2}{\sqrt{\alpha}}$	$\frac{6}{\alpha} + 3$
Beta(α, β)	$\frac{\alpha}{\alpha+\beta}$	(no closed form)	$\frac{\alpha-1}{\alpha+\beta-2}$	$\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$	$\frac{2(\beta-\alpha)\sqrt{\alpha+\beta+1}}{(\alpha+\beta+2)\sqrt{\alpha\beta}}$	*

$$* \text{Kurt}(\text{Beta}(\alpha, \beta)) = \frac{6[(\alpha-\beta)^2(\alpha+\beta+1) - \alpha\beta(\alpha+\beta+2)]}{\alpha\beta(\alpha+\beta+2)(\alpha+\beta+3)} + 3$$

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Sample Moments

Suppose

$$X_1, \dots, X_n$$

are i.i.d. random variables.

Definition

The k th sample moment is

$$M_k = \frac{1}{n} \sum_{j=1}^n X_j^k.$$

- ▶ Sample moments estimate population moments.
- ▶ The first sample moment, $M_1 = \frac{1}{n} \sum_{j=1}^n X_j$, is an important quantity in Statistics, also known as the *Sample Mean*, denoted by $\bar{X}$. It estimates the first population moment, i.e., $\mu'_1 = \mathbb{E}[X]$, the population mean.

i.i.d. Random Variables (Random Samples)

Definition: i.i.d. Random Variables

We say that the random variables $X_1, \dots, X_n$ are *independent and identically distributed* (i.i.d.) with distribution F , and write

$$X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} F,$$

if:

- ▶ $X_1, \dots, X_n$ are independent;
- ▶ each X_j has distribution F .

Terminology: If $X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} F$, then we also call a *random sample* of size n from the population distribution F . Quantities associated with F are called *population quantities*, while quantities computed from the random sample are called *sample quantities*.

Example

If

$$X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} N(\mu, \sigma^2),$$

then for every $j = 1, \dots, n$,

$$\mathbb{E}(X_j) = \mu, \quad \text{Var}(X_j) = \sigma^2.$$

Example of a Random Sample

Suppose the population consists of the GPAs of all University of Washington undergraduate students. Let F denote the population distribution of undergraduate GPAs.

Now suppose we randomly select 12 undergraduate students and record their GPAs:

$$X_1, \dots, X_{12} \stackrel{\text{i.i.d.}}{\sim} F.$$

Here:

- ▶ each X_j represents one student's GPA;
- ▶ all X_j 's have distribution F ;
- ▶ the observations are assumed to be independent.

A possible sample is:

3.5, 3.7, 4.0, 2.9, 3.1, 3.8, 3.4, 2.7, 3.9, 3.2, 3.6, 3.3.

Important

In practice, we usually do not know the population distribution F or its parameters, such as the population mean μ and variance σ^2 .

Instead, we observe a random sample

$$X_1, \dots, X_n$$

from the population and use the random sample to estimate unknown population quantities.

Sample Mean and Sample Variance

Let

$$X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} F, \quad \mathbb{E}(X_j) = \mu, \quad \text{Var}(X_j) = \sigma^2.$$

Definitions

The *sample mean* and *sample variance* are

$$\bar{X}_n = \frac{1}{n} \sum_{j=1}^n X_j, \quad S_n^2 = \frac{1}{n-1} \sum_{j=1}^n (X_j - \bar{X}_n)^2.$$

Mean and Variance

$$\mathbb{E}(\bar{X}_n) = \mu, \quad \text{Var}(\bar{X}_n) = \frac{\sigma^2}{n}.$$

$$\mathbb{E}(S_n^2) = \sigma^2, \quad \text{Var}(S_n^2) = \frac{1}{n} \left(\mu_4 - \frac{n-3}{n-1} \sigma^4 \right),$$

where

$$\mu_4 = \mathbb{E}[(X_j - \mu)^4].$$

Summary

- ▶ Mean, median, and mode describe center differently.
- ▶ Higher moments describe asymmetry and tail behavior.
- ▶ Skewness measures asymmetry.
- ▶ Kurtosis measures tail heaviness.
- ▶ Sample moments estimate population moments.
- ▶ The variance of the sample mean decreases like

$$\frac{1}{n}.$$

Next lecture

Moment generating functions.